Does regional variation in wage levels identify the effects of a national minimum wage?

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NBER Labor Studies Program Meeting Fall 2023

Identifying minimum wage effects



Minimum wage increases in blue states

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Identifying minimum wage effects



Minimum wage increases in blue states

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Identifying minimum wage effects



$$y_{r,t} = \alpha_r + \delta_t + \beta T_r \cdot \mathbf{1} \{ t \ge 1 \} + \epsilon_{r,t}$$

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A national change in the minimum wage?



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A national change in the minimum wage?



 $y_{r,t} = \alpha_r + \delta_t + \beta T_r \cdot \mathbf{1} \{ t \ge 1 \} + \epsilon_{r,t}$

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The fraction affected design



$$y_{r,t} = \alpha_r + \delta_t + \beta FA_r \cdot \mathbf{1} \{ t \ge 1 \} + \epsilon_{r,t}$$

• Card (1992)

• Bailey, DiNardo and Stuart (2021); Dustmann et al. (2021, using *Gap*)

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The effective minimum wage design



 $y_{r,t} = \alpha_r + \delta_t + \beta \left[m w_t - w_{50,r,t} \right] + \gamma \left[m w_t - w_{50,r,t} \right]^2 + \epsilon_{r,t}$

• Lee (1999); older stuff like Neumark and Wascher (1992)

• Bosch and Manacorda (2010); Engbom and Moser (2022)

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This paper

Does regional variation in wage levels identify the effects of a national minimum wage?

• Evaluate identification assumptions of both designs

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This paper

Does regional variation in wage levels identify the effects of a national minimum wage?

• Evaluate identification assumptions of both designs

Findings for the **effective minimum wage design:**

- Identification assumptions of Lee (1999) are crucial, but hard to satisfy without region-specific minimum wages
- If regional variation is available: IV strategies a la Autor, Manning, and Smith (2016) are preferable

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This paper

Does regional variation in wage levels identify the effects of a national minimum wage?

• Evaluate identification assumptions of both designs

Findings for the **fraction affected/gap designs:**

- Parallel trend violations can come from unexpected places like e.g., skill-biased technical change
- Pre-trends checks useful, but should be implemented with care
- May be sensitive to functional form assumptions

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Outline

Setup

- e Effective minimum wage design
- Fraction affected/gap designs
- Onclusion

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The data-generating process

$$\mathbf{y}_{r,t} = f\left(mw_t, \boldsymbol{\theta}_{r,t}\right)$$

- $r \in \{1, \dots, R\}$ are regions, $t \in \{0, 1\}$ is time
- **y**_{*r*,*t*}: vector of outcomes, e.g., employment to population ratio and quantiles of the log wage distribution
- f given by an economic model
- $mw_1 > mw_0$
- $\theta_{r,t}$: Region-time-specific parameters

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The data-generating process

$$\mathbf{y}_{r,t} = f\left(mw_t, \boldsymbol{\theta}_{r,t}\right)$$

What this rules out:

- Regional spillovers due to e.g. migration: Cadena (2014)
- Short vs. long effects of minimum wages: Sorkin (2015), Hurst et al. (2022), Vogel (2023)
- Bias caused by measurement error in region-level statistics: Autor, Manning, and Smith (2016)
- Diff-in-diffs with staggered treatment: de Chaisemartin and D'Haultfoeuille (2020)

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The causal effect of interest

$$\mathbf{y}_{r,t} = f\left(mw_t, \boldsymbol{\theta}_{r,t}\right)$$

$$\begin{aligned} \mathbf{ATE}_{0} &= \frac{1}{R} \sum_{r} f\left(mw_{1}, \theta_{r,0}\right) - \mathbf{y}_{r,0} \\ \mathbf{ATE}_{1} &= \frac{1}{R} \sum_{r} \mathbf{y}_{r,1} - f\left(mw_{0}, \theta_{r,1}\right) \\ \mathbf{ATE} &= \frac{\mathbf{ATE}_{0} + \mathbf{ATE}_{1}}{2} \end{aligned}$$

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Effective minimum wage design

Baseline specification for measuring wage spillover effects:

$$w_{q,r,t} - w_{0.5,r,t} = \alpha_{q,r} + \delta_{q,t} + \beta_q \left[mw_t - w_{0.5,r,t} \right] \\ + \gamma_q \left[mw_t - w_{0.5,r,t} \right]^2 + \epsilon_{q,r,t}$$

- $w_{q,r,t}$: quantile q of log wage distribution in r, t
- Each q is a separate regression Figure from AMS

$$\widehat{ATE}_{q} = \frac{1}{R} \sum_{r} \left\{ \hat{\beta}_{q} \left[(mw_{1} - w_{0.5,r,1}) - (mw_{0} - w_{0.5,r,0}) \right] + \hat{\gamma}_{q} \left[(mw_{1} - w_{0.5,r,1})^{2} - (mw_{0} - w_{0.5,r,0})^{2} \right] \right\}$$

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Also used for employment to population ratio (e.g., Engbom and Moser, 2022):

$$emp_{r,t} = \alpha_r + \delta_t + \beta \left[mw_t - w_{0.5,r,t}\right] + \gamma \left[mw_t - w_{0.5,r,t}\right]^2 + \epsilon_{r,t}$$

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Lee's (1999) semiparametric model

CDF of *latent* log wages has the form:

$$F_t\left(rac{w-\mu_{r,t}}{\sigma_{r,t}}
ight)$$
 with $F_t(0)=0.5$

- $\mu_{r,t}$ is the location (or *centrality*) parameter
- $\sigma_{r,t}$ is the *dispersion* parameter

Location (centrality) and dispersion



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CDF of *latent* log wages has the form:

$$F_t\left(rac{w-\mu_{r,t}}{\sigma_{r,t}}
ight)$$
 with $F_t(0)=0.5$

• Assumption 1: $w_{0.5,r,t} \approx \mu_{r,t}$

• Assumption 2: $\mu_{r,t}$ and $\sigma_{r,t}$ are uncorrelated conditional on t

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Building intuition with a linear model

The economic model is:

$$y_{q,r,t} = \alpha_{q,r} + \delta_{q,t} + \beta_q \left[m w_t - \mu_{r,t} \right] + \epsilon_{q,r,t}$$

We're interested in β_q . Taking differences:

$$y_{q,r,1} - y_{q,r,0} = (\delta_{q,1} - \delta_{q,0}) + \beta_q \left[(mw_1 - \mu_{r,1}) - (mw_0 - \mu_{r,0}) \right] \\ + (\epsilon_{q,r,1} - \epsilon_{q,r,0}) \\ = \rho_q + \beta_q \left[-(\mu_{r,1} - \mu_{r,0}) \right] + \upsilon_{q,r}$$

Key source of variation: shocks to location parameter.

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The ideal scenario: A is treated, B is not



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The ideal scenario: A is treated, B is not



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Alternative scenario: no idiosyncratic shocks



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If we could observe $\mu_{r,t}$: no variation



Economic model: $\Delta y_{q,r} = \rho_q + \beta_q \left[-(\mu_{r,1} - \mu_{r,0}) \right] + \upsilon_{q,r}$

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Min. wage effects on median: bad variation



Statistical model: $\Delta y_{q,r} = \rho_q + \beta_q \left[-(w_{0.5,r,1} - w_{0.5,r,0}) \right] + v_{q,r}$

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Issue #1: Correlated measurement error

- Causal effects on median wage likely to be correlated with effects on other quantiles and on employment
- Problem exists even if average effect on median wage is zero
- Is it quantitatively relevant?

Simulations: the Normal-markdown model

- Latent log wages are Normal
- Vector $\theta_r = [\mu_{r,0}, \sigma_{r,0}, \mu_{r,1}, \sigma_{r,1}]$ drawn from multivariate Normal calibrated to match data from the US CPS
 - Regions are states
 - Years are 1989 and 2004
 - Construct different scenarios. For example:
 - Shut down differences in dispersion: $\sigma_{r,t} = \bar{\sigma}_t$
 - Allow dispersion in $\sigma_{r,t}$, but make it uncorrelated with $\mu_{r,t}$

Minimum wage in the Normal-markdown model



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Panel A has $\sigma_{r,t} = \bar{\sigma}_t$ and $Corr(\mu_{r,0}, \mu_{r,1}) = 0.89$

	Outcome					
	Emp.	o. p10 - p50 p25 - p50 p90 -				
Panel A: Regions differ only in location parameter						
True average causal effect	-0.010	0.019	0.006	-0.004		
Effective min. wage	-0.010	0.019	0.006	-0.004		
	(0.001)	(0.002)	(0.000)	(0.000)		

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Model imposes $Corr(\mu_{r,t}, \sigma_{r,t}) = 0.$

	Outcome					
	Emp.	p. p10 - p50 p25 - p50 p90 -				
Panel B: Regions differ in location and dispersion						
True average causal effect	-0.010	0.020	0.006	-0.004		
Effective min. wage	-0.007	0.033	0.014	-0.023		
	(0.004)	(0.023)	(0.013)	(0.028)		

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	Outcome						
	Emp. p10 - p50 p25 - p50 p90 - p5						
Panel C: As above, but larger increase in min. wage							
True average causal effect	-0.032	0.078	0.017	-0.012			
Effective min. wage	-0.013	0.115	0.045	-0.079			
	(0.015)	(0.040)	(0.023)	(0.053)			

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	Outcome					
	Emp. p10 - p50 p25 - p50 p90 -					
Panel D: St. dev. of dispersion is 50% larger						
True average causal effect	-0.010	0.020	0.006	-0.004		
Effective min. wage	-0.003	0.050	0.025	-0.047		
	(0.006)	(0.033)	(0.020)	(0.041)		

Issue #2: Correlation between $\mu_{r,t}$ and $\sigma_{r,t}$



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Issue #2: Correlation between $\mu_{r,t}$ and $\sigma_{r,t}$

	Outcome					
	Emp.	Emp. p10 - p50 p25 - p50 p90 -				
Panel B: Contemporaneous correlation of 0.076						
True average causal effect	-0.010	0.020	0.006	-0.004		
Effective min. wage	-0.002	0.076	0.040	-0.075		
	(0.004)	(0.021)	(0.012)	(0.026)		

Correlation between mean log wage and std. dev. of log wage at state level:

- 0.076 in 1989
- 0.264 in 2004

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Should we expect such correlations to occur?

- Regional differences in workforce composition
 - E.g., Lemieux (2006): there's more wage dispersion in more higher-wage education-experience cells
- Differences in endowments affecting industrial composition may also affect both wage levels and wage dispersion

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Can we fix the problem with appropriate controls?

- Options: control for worker composition, regional trends...
- Only helps if *residual variation* in $\mu_{r,t}$ is uncorrelated with $\sigma_{r,t}$.

Lee (1999) argues *fewer* controls can be better:

"... the reduced identifying variation resulting from eliminating the "permanent" state effects may magnify biases due to misspecification, in the same way biases stemming from measurement error in the independent variable are magnified when true variation in the independent variable is reduced."

 In simulations: estimator without region fixed effects works well for wage spillover effects, but not for employment.

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Other specifications and diagnostics

- Removing time fixed effects? No. Table
- Osing higher quantiles as the deflator? No. Table
- IV strategy by Autor, Manning, and Smith (2016)? Table
 - Yes, but only feasible with state-level minimum wages.
 - Version with two, rather than three instruments may be better.
- Oiagnosing with upper tail spillovers?
 - Subject to false positives and false negatives. Why?
- In the second second
 - No; paper includes exercises with the Canonical CES Model or the one from (Haanwinckel (2023)).

If you have exogenous variation in region-level minimum wages:

• Use instrumental variables approaches to isolate that variation

If you don't:

- Use the median as deflator and include time fixed effects
- The argument for identification should be:
 - Is there a structural shock that shifts location $\mu_{r,t}$, but not dispersion $\sigma_{r,t}$ or latent employment (conditional on controls?)
 - Is that good variation large enough to offset biases caused by the imperfect measurement of location?

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Fraction affected and Gap designs

Definition: Fraction Affected

Fraction Affected Illustration > Definition: Gap measure

Issues:

- Misspecification biases **Details**
- Regression to the mean **Details** 2
- Ommon trends in the dispersion of latent wages

Outline

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Contribution: pointing out potential pitfalls.

What if there is no regional variation in minimum wage laws, and neither of those between-region designs are recommended?

- Within-region comparisons of affected vs. non-affected firms or workers
- Structural approaches
- Thanks!

OLS estimates for the US, 1972-2012



Panel C. Males and females—state fixed effects and trends.

Figure 3. OLS Estimates of the Relationship between log(p) - log(p50) and log(min) - log(p50)AND ITS SOUARE, 1979–2012

• From Autor, Manning and Smith (2016)

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Removing time fixed effects

		Ou	tcome		
	Emp.	p10 - p50	p25 - p50	p90 - p50	
Panel A: Regions differ only in location, stable distribution					
True average causal effect	-0.010	0.019	0.006	-0.004	
Effective min. wage	-0.010	0.020	0.006	-0.004	
	(0.001)	(0.003)	(0.000)	(0.000)	
Eff. min. wage, no time FE	-0.010	0.019	0.006	-0.004	
	(0.000)	(0.001)	(0.000)	(0.000)	
Panel B: Regions differ in locatio	n and disp	ersion			
True average causal effect	-0.010	0.020	0.006	-0.004	
Effective min. wage	-0.007	0.034	0.015	-0.023	
	(0.004)	(0.023)	(0.013)	(0.028)	
Eff. min. wage, no time FE	-0.007	0.052	0.024	-0.041	
	(0.001)	(0.006)	(0.003)	(0.007)	

Percentile 90 as the deflator

	Outcome					
	Emp.	p10 - p90	p25 - p90	p50 - p90		
Panel A: Regions differ only in location, stable distribution						
True average causal effect	-0.010	0.023	0.009	0.000		
Effective min. wage, p90	-0.010	0.024	0.009	0.000		
	(0.001)	(0.003)	(0.001)	(0.000)		
Panel B: Regions differ in locati	ion and dis	spersion				
True average causal effect	-0.010	0.023	0.009	0.000		
Effective min. wage, p90	0.009	0.219	0.176	0.000		
	(0.003)	(0.026)	(0.021)	(0.000)		
Panel C: Model from Haanwinc	kel (2023)					
True average causal effect	-0.046	0.211	0.090	0.000		
Effective min. wage, p90	0.025	0.384	0.306	0.000		
	(0.012)	(0.015)	(0.021)	(0.000)		

State-level variation and IV approaches

	Outcome				
	Emp.	p10 - p50	p25 - p50	p90 - p50	
Panel A: No regional variation in minimum wage.					
True average causal effect	-0.010	0.020	0.006	-0.004	
Effective min. wage	-0.002	0.076	0.040	-0.075	
	(0.004)	(0.021)	(0.012)	(0.026)	
Panel B: 20% of regions with lo	ocal min. w	vage			
True average causal effect	-0.015	0.035	0.008	-0.005	
Effective min. wage	-0.015	0.050	0.014	-0.018	
	(0.003)	(0.009)	(0.005)	(0.012)	
Two instruments	-0.016	0.036	0.009	-0.006	
	(0.004)	(0.013)	(0.008)	(0.017)	
Three instruments (AMS)	-0.017	0.041	0.008	-0.005	
	(0.003)	(0.010)	(0.006)	(0.013)	

Diagnosing with upper tail spillovers

Subject to false positives:

 Models such as Engbom and Moser (2023) and Haanwinckel (2023) have mechanisms that generate spillovers high into the wage distribution, such as reallocation to higher-wage firms (as in Dustmann et al. 2021)

Subject to false negatives:

 Estimator may be biased for employment and lower-tail spillovers, but unbiased for upper-tail spillovers (especially after combining effects of issues discussed here with measurement error-induced bias from Autor, Manning, and Smith 2016).

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Canonical model **Grack**

	Outcome					
	Emp.	p10 - p50	p25 - p50	p90 - p50		
Panel A: Initial minimum wage is le	ow, sigma	= 3.0				
True average causal effect	-0.009	0.017	0.005	-0.004		
Eff. min. wage, no region FE	-0.009	0.018	0.003	0.044		
	(0.000)	(0.002)	(0.001)	(0.003)		
Panel B: Initial minimum wage is lo	ow, sigma	= 1.4				
True average causal effect	-0.009	0.016	0.005	-0.003		
Eff. min. wage, no region FE	-0.010	0.002	-0.007	0.078		
	(0.000)	(0.003)	(0.002)	(0.003)		
Panel D: Initial minimum wage is h	nigh, sigma	a = 1.4				
True average causal effect	-0.039	0.054	0.019	-0.016		
Eff. min. wage, no region FE	-0.045	0.042	0.008	0.067		
	(0.001)	(0.002)	(0.002)	(0.003)		

Model from Haanwinckel (2023)

	Outcome					
	Emp.	p10 - p50	p25 - p50	p90 - p50		
True average causal effect	-0.046	0.198	0.077	-0.013		
Effective min. wage	-0.015	0.218	0.122	0.070		
	(0.012)	(0.011)	(0.013)	(0.037)		
Effective min. wage, no region FE	-0.073	0.196	0.088	-0.016		
	(0.006)	(0.005)	(0.006)	(0.014)		
Effective min. wage, no time FE	0.113	0.212	0.121	-0.139		
	(0.004)	(0.003)	(0.004)	(0.012)		
AMS, no time FE	0.125	0.211	0.121	-0.159		
	(0.005)	(0.003)	(0.003)	(0.011)		

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Fraction affected design

$$y_{o,r,t} = \alpha_{o,r} + \delta_{o,t} + \beta_o FA_r \cdot \mathbf{1}\{t = 1\} + \epsilon_{o,r,t}$$

- Each outcome o is a separate regression
- FA_r : share of workers *i* with $mw_0 \le w_i < mw_1$ at time t = 0
- $A\hat{T}E_o$ is the product of average FA and β_o
- Typical application: one-time mw hike following years of stability
 - Can test for differential pre-trends

Fraction affected design ••••



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Gap design

$$y_{o,r,t} = \alpha_{o,r} + \delta_{o,t} + \beta_o Gap_r \cdot \mathbf{1}\{t = 1\} + \epsilon_{o,r,t}$$
$$Gap_r = \frac{\sum_{i=1}^{l_r} \max\{\exp(mw_1) - \exp(w_{i,0}), 0\}}{\sum_{i=1}^{l_r} \exp(w_{i,0})}$$

• Introduced by Card and Krueger (1994) at the firm level; later used at the regional level (e.g., Dustmann et al., 2021)

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Issue #1: misspecification biases \blacksquare

- Fraction affected and gap designs define "susceptibility to treatment" based on a theoretically-inspired measure
 - But not the reduced form of some popular economic model
- Simulations show biases even in "ideal" applications
 - Attenuation seems more prevalent for employment and wage effects in the lower tail, when the minimum wage is strongly binding

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Normal-markdown model, 1/2 (Back)

Regions differ only in time-invariant location parameter μ_r

			Outcome				
	Emp.	p10	p25	p50	p90		
Panel A: Initial min. wage is small, $m = 0.7$							
True average causal effect	-0.006	0.016	0.008	0.004	0.002		
Fraction affected	-0.008	0.020	0.010	0.005	0.003		
	(0.000)	(0.002)	(0.001)	(0.002)	(0.002)		
Gap measure	-0.006	0.015	0.007	0.004	0.002		
	(0.000)	(0.001)	(0.001)	(0.001)	(0.001)		
Panel B: Initial min. wage is sm	all, $m = 0.$	9					
True average causal effect	-0.018	0.043	0.022	0.012	0.006		
Fraction affected	-0.019	0.038	0.022	0.013	0.006		
	(0.000)	(0.001)	(0.001)	(0.002)	(0.002)		
Gap measure	-0.015	0.029	0.017	0.010	0.005		
	(0.000)	(0.002)	(0.001)	(0.001)	(0.001)		

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Normal-markdown model, 2/2

Regions differ only in time-invariant location parameter μ_r

			Outcome				
	Emp.	p10	p25	p50	p90		
Panel C: Initial min. wage is large, $m = 0.7$							
True average causal effect	-0.042	0.162	0.049	0.028	0.013		
Fraction affected	-0.053	0.153	0.063	0.036	0.018		
	(0.001)	(0.021)	(0.005)	(0.002)	(0.002)		
Gap measure	-0.037	0.101	0.044	0.025	0.013		
	(0.001)	(0.017)	(0.003)	(0.002)	(0.002)		
Panel D: Initial min. wage is lar	ge, $m = 0.9$	9					
True average causal effect	-0.079	0.126	0.084	0.053	0.027		
Fraction affected	-0.073	0.071	0.067	0.052	0.030		
	(0.002)	(0.003)	(0.002)	(0.002)	(0.002)		
Gap measure	-0.052	0.050	0.048	0.037	0.021		
	(0.002)	(0.003)	(0.002)	(0.002)	(0.002)		

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Issue #2: Regression to the mean \blacksquare

- Treated regions selected on extreme outcomes
 - Sampling variation (if low-wage regions have small samples)
 - Regional TFP (Caliendo et al. 2017; Gennaioli et al. 2014)
- May be captured by tests for pre-trends...
 - ... but not if treatment is defined by averaging the fraction affected over all pre-treatment years
- Controlling for a linear trend does not help

Normal-markdown model

	Outcome								
	Emp.	p10	p25	p50	p90				
Panel A: Only permanent differences in location									
True average causal effect	-0.010	0.026	0.012	0.007	0.003				
Gap measure	-0.009	0.027	0.011	0.006	0.003				
	(0.000)	(0.003)	(0.001)	(0.001)	(0.001)				
Panel B: Adding location shocks, stable distributions									
True average causal effect	-0.010	0.026	0.012	0.007	0.003				
Gap measure	-0.007	0.043	0.031	0.027	0.024				
	(0.001)	(0.010)	(0.012)	(0.012)	(0.013)				
Panel C: Adding dispersion differences and shocks, stable distributions									
True average causal effect	-0.010	0.026	0.013	0.007	0.003				
Gap measure	-0.007	0.052	0.034	0.024	0.008				
	(0.002)	(0.012)	(0.012)	(0.012)	(0.018)				

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Issue #3: Trends in dispersion of latent wages \blacksquare



Daniel Haanwinckel (UCLA, NBER)

NBER Labor Studies Fall 2023

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Issue #3: Trends in dispersion of latent wages \blacksquare



Daniel Haanwinckel (UCLA, NBER)

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Normal-markdown model

	Outcome							
	Emp.	p10	p25	p50	p90			
Panel C: Adding dispersion differences and shocks, stable distributions								
True average causal effect	-0.010	0.026	0.013	0.007	0.003			
Gap measure	-0.007	0.052	0.034	0.024	0.008			
	(0.002)	(0.012)	(0.012)	(0.012)	(0.018)			
Panel D: Average dispersion falls over time								
True average causal effect	-0.010	0.026	0.013	0.007	0.003			
Gap measure	-0.004	0.046	0.032	0.023	0.009			
	(0.002)	(0.013)	(0.013)	(0.013)	(0.019)			

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