

# ECONOMETRICA

JOURNAL OF THE ECONOMETRIC SOCIETY

*An International Society for the Advancement of Economic  
Theory in its Relation to Statistics and Mathematics*

<http://www.econometricsociety.org/>

*Econometrica*, Vol. 78, No. 5 (September, 2010), 1569–1591

## NONPARAMETRIC IDENTIFICATION AND ESTIMATION OF NONADDITIVE HEDONIC MODELS

JAMES J. HECKMAN

*University of Chicago, Chicago, IL 60637, U.S.A., and University College Dublin, American Bar  
Foundation and Cowles Foundation, Yale University*

ROSA L. MATZKIN

*University of California, Los Angeles, Los Angeles, CA 90095, U.S.A.*

LARS NESHEIM

*University College London, London, WC1E 6BT, U.K. and CEMMAP and Institute for Fiscal  
Studies*

---

The copyright to this Article is held by the Econometric Society. It may be downloaded, printed and reproduced only for educational or research purposes, including use in course packs. No downloading or copying may be done for any commercial purpose without the explicit permission of the Econometric Society. For such commercial purposes contact the Office of the Econometric Society (contact information may be found at the website <http://www.econometricsociety.org> or in the back cover of *Econometrica*). This statement must be included on all copies of this Article that are made available electronically or in any other format.

---

## NONPARAMETRIC IDENTIFICATION AND ESTIMATION OF NONADDITIVE HEDONIC MODELS

BY JAMES J. HECKMAN, ROSA L. MATZKIN, AND LARS NESHEIM<sup>1</sup>

This paper studies the identification and estimation of preferences and technologies in equilibrium hedonic models. In it, we identify nonparametric structural relationships with nonadditive heterogeneity. We determine what features of hedonic models can be identified from equilibrium observations in a single market under weak assumptions about the available information. We then consider use of additional information about structural functions and heterogeneity distributions. Separability conditions facilitate identification of consumer marginal utility and firm marginal product functions. We also consider how identification is facilitated using multimarket data.

KEYWORDS: Hedonic models, hedonic equilibrium, nonadditive models, identification, nonparametric estimation.

### 1. INTRODUCTION

HOW MUCH ARE CONSUMERS WILLING TO PAY for increased product quality? What wage premium do workers require to compensate them for an increase in on-the-job fatality risk? How much are homeowners willing to pay to avoid moving one mile further from work? What will be the effect of new clean air legislation on real estate prices? Hedonic models provide information that helps to narrow the range of possible answers to these questions. In an hedonic market, the price function for a good or service decomposed into its attributes measures the equilibrium relationship between attributes of the good and the price at which buyers and sellers are willing to trade. It is the price that equates supply and demand in the space of hedonic attributes. It provides an exact measure of the marginal willingness to pay and willingness to accept for equilibrium transactions in a market. However, in general, the equilibrium price function does not identify the willingness to pay for counterfactual transactions that might occur in markets characterized by different endowments, technology, and preferences. Due to unobserved heterogeneity,

<sup>1</sup>This research was supported by NSF Grants SES-024158 and BCS-0433990/BCS-0852261. Lars Nesheim also thanks the Leverhulme Trust and the U.K. Economic and Social Research Council (ESRC Grant RES-589-28-0001) for support through its funding of the Centre for Micro-data Methods and Practice (<http://www.cemmap.ac.uk>). Rosa L. Matzkin's research has been supported also by NSF Grant SES-0551272/SES-0833058. We thank the co-editor, Whitney Newey, and three referees for their extremely useful input and the participants at seminars at Northwestern University, Universidad de San Andres, Princeton University, Harvard/MIT, UCLA/USC, University of California at Berkeley, University of Minnesota, the Bureau of Labor Statistics, University of Chicago, Universidad Torcuato Di Tella, the 2002 Workshop on Characteristics Models: Theory and Applications (University of Copenhagen), the 2002 Workshop on Mathematical Economics (IMPA), and the 2004 Banff International Research Station (BIRS) workshop on "Mathematical Structures in Economic Theory and Econometrics" for their useful comments. We thank Myrna Wooders and Daniel McFadden for many stimulating conversations, Donald J. Brown for helpful comments, and Yong Hyeon Yang for excellent research assistance.

there is a selection problem.<sup>2</sup> To take a labor market example, workers who choose to work in jobs with a low fatality risk are likely to have low unobserved tolerance for risk. As a result, the compensation they require to accept more risk is likely to be higher than the equilibrium wage premium for risky jobs. To identify willingness to pay and willingness to accept for counterfactual transactions, one needs to identify the structural parameters of buyers and sellers.

Knowledge of the hedonic price function alone is not enough to analyze general equilibrium effects of policy changes in hedonic markets. When supply or demand conditions change, buyers and sellers will, in general, alter how they sort in hedonic markets. Hence, to analyze general equilibrium changes in hedonic markets, it is necessary to identify the structural parameters.<sup>3</sup>

Observed and unobserved heterogeneity play key roles in hedonic equilibrium sorting models. In particular, their distributions and the underlying utility and production functions of the agents generate equilibrium sorting outcomes and shape how an hedonic economy will react to changes in the economic environment. Previous studies of hedonic markets have analyzed either parametric or restrictive nonparametric utility and production functions, where the unobserved heterogeneity is specified as a term added to marginal utility or marginal product functions.<sup>4</sup> Ekeland, Heckman, and Nesheim (2004) analyzed nonparametric hedonic models with additive marginal utility and additive marginal product functions. They showed that hedonic models with additivity restrictions are nonparametrically identified with single market data, and they presented two methods for recovering the structural functions in such models.

The additivity assumptions used to establish identification in Ekeland, Heckman, and Nesheim (2004) are strong. No heterogeneity in the curvature of production and preference functions is allowed. In addition, these restrictions are overidentifying.

In contrast, this paper establishes nonparametric identification of structural functions and distributions in general hedonic models without imposing additivity. We allow the curvature of the marginal utility for the product attribute, as well as the distribution of the marginal utilities, to vary in general ways across agents with different observed characteristics. We allow an analogous property to hold for nonadditive unobserved heterogeneity in marginal product functions on the supply side of the market.

We first analyze what can be identified in an unrestricted model fit using data from a single market that satisfies standard economic regularity conditions and additional mild statistical regularity conditions. We then restrict the functional structure of the marginal utilities and marginal product functions,

<sup>2</sup>See, for example, Rosen (1974).

<sup>3</sup>See, for example, Hurwicz (1962) or, for a more modern statement, Heckman and Vytlačil (2007).

<sup>4</sup>For analyses of identification of parametric hedonic models, see Brown and Rosen (1982), Brown (1983), Epplé (1987), Bajari and Benkard (2005), or Berry and Pakes (2007).

obtaining nonparametric versions of random coefficient and household scale models. We show that in these cases, the marginal utility and marginal product functions, as well as the distribution of the unobserved characteristics of the agents, are identified. Our proofs of identification are constructive and suggest nonparametric estimation strategies whose properties we examine in the Technical Appendix (Heckman, Matzkin, and Nesheim (2010)).<sup>5</sup>

We go on to show that models that cannot be identified in a single cross section can be identified using data from multiple markets. This follows from the economics of the hedonic model. In general, variation in the distributions of observed variables across markets induces cross-market variation in price functions. This part of our analysis formalizes and extends discussions in Rosen (1974), Brown and Rosen (1982), Epple (1987), and Kahn and Lang (1988).

To focus on the key identification problems that arise in a setup where marginal utilities are nonadditive in unobserved heterogeneity, we concentrate on an hedonic model with a single product characteristic. Our main results can be extended to hedonic models with multiple characteristics captured by an index structure, as in Epple and Sieg (1999) and Sieg, Smith, Banzhaf, and Walsh (2004). Our methodology extends the two stage method of Rosen (1974) to a nonparametric setting. The first stage estimates the hedonic price function and its derivatives. Using the estimated price function together with the first order conditions of the agents identifies marginal production and utility functions.

This paper proceeds in the following way. Section 2 describes the hedonic model for a product with a single attribute. Section 3 studies the identification of nonadditive marginal utility and nonadditive marginal product functions. Section 4 analyzes identification in multiple markets. Nonparametric estimators are discussed in Section 5. Their properties are derived in the Technical Appendix (see Heckman, Matzkin, and Nesheim (2010)). The Technical Appendix also presents a limited Monte Carlo analysis. Section 6 summarizes and discusses possible extensions.

## 2. THE COMPETITIVE HEDONIC EQUILIBRIUM MODEL

The model analyzed in this paper applies to a spot market in which products are differentiated by their attributes, prices are set competitively, and participating buyers and sellers each trade a single type of product chosen from a set of feasible products. Our analysis of demand in this section and in Section 3 can be applied to hedonic demand estimation problems in imperfectly competitive markets with a continuum of potential products where consumers are fully informed price takers. The multimarket analysis in Section 4 also can readily be adapted to study demand in imperfectly competitive markets with a continuum of products as long as there is sufficient cross-market variation in market structure to induce price variation. We focus on competitive markets for ease of exposition.

<sup>5</sup>See Technical Appendices D and F.

To fix ideas, we consider a labor market setting in which jobs are characterized by their attributes. Workers (sellers) match to single employee firms (buyers) and supply job services indexed by a scalar  $z$ , where  $z \in \tilde{Z}$ .<sup>6</sup> We assume that  $\tilde{Z} \subseteq \mathbf{R}$  is compact. The space  $\tilde{Z}$  is the space of technologically feasible job attributes. The variable  $z$  denotes a job attribute assumed to be a disamenity for the workers and an input for the firms. For example,  $z$  could measure the probability of injury on the job as in [Kniesner and Leeth \(1995\)](#) or [Leeth and Ruser \(2003\)](#), so  $\tilde{Z} = [0, 1]$ . Let  $P(z)$  be a twice continuously differentiable price function defined on  $\tilde{Z}$ . The value of  $P(z)$  is the wage paid at a job characterized by attribute  $z$ .

Each worker has quasilinear utility function  $P(z) - U(z, x, \varepsilon)$ , where  $x$  is a vector of observed characteristics of dimension  $n_x$  and  $\varepsilon$  is a scalar unobserved heterogeneity term.<sup>7</sup> Observe that  $U(\cdot)$  is a disutility.  $\varepsilon$  is statistically independent of  $x$ . In the risk of injury example,  $\varepsilon$  is interpreted as the component of unobserved preference for risk that is independent of observables such as age, education, gender, race, and so forth. Note that since  $\varepsilon$  is allowed to enter marginal utility  $U_z$  in a nonadditive way, the independence assumption is much weaker than it would be if  $\varepsilon$  entered additively. In particular, the nonadditive specification is consistent with a model where  $U_z$  is a function of an unobserved  $\xi = m(x, \varepsilon)$ , whose distribution depends on the vector of observed characteristics,  $x$ , of the agent. The population of workers is described by the pair of density functions  $f_x$  and  $f_\varepsilon$ , which are assumed to be differentiable and strictly positive on the compact sets  $\tilde{X} \subseteq \mathbf{R}^{n_x}$  and  $\tilde{E} \subseteq \mathbf{R}$ , respectively. Each worker may choose not to trade, in which case they obtain reservation utility  $V_0$ .

Each firm has a production function  $\Gamma(z, y, \eta)$ , where  $y$  is a vector of observed characteristics of the firm of dimension  $n_y$  and  $\eta$  is a scalar unobserved heterogeneity term. We assume that  $\eta$  is statistically independent of  $y$  and that  $(y, \eta)$  is independent of  $(x, \varepsilon)$ . The population of firms is described by the pair of density functions  $f_y$  and  $f_\eta$  that are strictly positive on the compact sets  $\tilde{Y} \subseteq \mathbf{R}^{n_y}$  and  $\tilde{H} \subseteq \mathbf{R}$ , respectively. If a firm opts out of the market, it earns reservation profits  $\Pi_0$ .

Both  $U$  and  $\Gamma$  are assumed to be twice continuously differentiable with respect to all arguments.<sup>8</sup> Additionally, we assume that, given the equilibrium price function, each worker and firm that participates has a unique interior op-

<sup>6</sup>If jobs are characterized by a vector of attributes  $a \in \tilde{A} \subseteq \mathbf{R}_n$ , we assume that preferences and costs depend on  $a$  only through the scalar index  $z = g(a)$ . We discuss identification of a single index model in Section 3. Alternatively, if utility is quasilinear and both utility and production functions are additively separable functions of  $z$  and other attributes, the price function will share these properties and the market for  $z$  can be analyzed independently of the other attributes.

<sup>7</sup>When utility is not quasilinear, we can identify the ratio of the marginal utility of  $z$  to the marginal utility of income (see Technical Appendix B).

<sup>8</sup>The equilibrium analysis can be extended to cases in which  $z$  is a vector, the arguments of the functions are discrete, the measures on characteristics are not absolutely continuous with respect

timum. A sufficient condition for this to hold is a Spence–Mirrlees type single-crossing condition, requiring that  $U_{zz} - \Gamma_{zz} > 0$ ,  $U_{z\varepsilon} < 0$ , and  $\Gamma_{z\eta} > 0$ , for all  $(z, x, \varepsilon, y, \eta)$ , be satisfied.<sup>9</sup> For simplicity, we assume that these conditions are satisfied.

Each worker chooses  $z \in \tilde{Z}$ , a job type or a location in the space of job attributes, to maximize

$$P(z) - U(z, x, \varepsilon).$$

Under the above conditions, it follows from Gale and Nikaido (1965) that there exists a supply function  $z = s(x, \varepsilon)$  such that

$$(2.1) \quad P_z(s(x, \varepsilon)) - U_z(s(x, \varepsilon), x, \varepsilon) = 0.$$

By the Implicit Function Theorem,

$$(2.2) \quad \frac{\partial s(x, \varepsilon)}{\partial \varepsilon} = \frac{U_{z\varepsilon}(s(x, \varepsilon), x, \varepsilon)}{P_{zz}(s(x, \varepsilon)) - U_{zz}(s(x, \varepsilon), x, \varepsilon)}$$

so that  $\frac{\partial s(x, \varepsilon)}{\partial \varepsilon} > 0$  since  $U_{z\varepsilon} < 0$  and the denominator is negative from the assumption of an interior optimum. Our assumptions imply that  $s(x, \varepsilon)$  is strictly monotonic in  $\varepsilon$  and that its inverse with respect to  $\varepsilon$  exists. We denote the inverse supply function by  $\varepsilon = \tilde{s}(z, x)$ .

A parallel analysis can be performed for the demand side of the market. Each firm chooses  $z \in \tilde{Z}$  to maximize the profit function

$$\Gamma(z, y, \eta) - P(z).$$

The resulting demand function,  $z = d(y, \eta)$ , satisfies

$$\Gamma_z(d(y, \eta), y, \eta) - P_z(d(y, \eta)) = 0,$$

and

$$\frac{\partial d(y, \eta)}{\partial \eta} = \frac{\Gamma_{z\eta}(d(y, \eta), y, \eta)}{P_{zz}(d(y, \eta)) - \Gamma_{zz}(d(y, \eta), y, \eta)}$$

so that  $\frac{\partial d(y, \eta)}{\partial \eta} > 0$  since  $\Gamma_{z\eta} > 0$  and the denominator is positive. We denote the inverse demand function  $\eta = \tilde{d}(z, y)$ .

The inverse supply and demand functions characterize equilibrium.<sup>10</sup> Let

$$Z_s = \{z \in \tilde{Z} \mid z = s(x, \varepsilon) \text{ for some } (x, \varepsilon) \in \tilde{X} \times \tilde{E}\}$$

---

to Lebesgue measure, or the functions are not differentiable. See Ekeland (2010) or Chiappori, McCann, and Nesheim (2010) for details.

<sup>9</sup>See Chiappori, McCann, and Nesheim (2010) for details.

<sup>10</sup>See Ekeland, Heckman, and Nesheim (2004).

be the range of the mapping  $s(x, \varepsilon)$ . Using a standard change of variables formula, the density of supply implied by the inverse supply function is

$$(2.3) \quad \int_{\tilde{X}} f_{\varepsilon}(\tilde{s}(z, x)) f_x(x) \frac{\partial \tilde{s}(z, x)}{\partial z} dx$$

for  $z \in Z_s$ . For  $z \in \tilde{Z} \setminus Z_s$ , the density of supply is zero.

Analogously, let

$$Z_d = \{z \in \tilde{Z} \mid z = d(y, \eta) \text{ for some } (y, \eta) \in \tilde{Y} \times \tilde{H}\}$$

be the range of the mapping  $d(y, \eta)$ . The density of the demanded  $z$  is

$$(2.4) \quad \int_{\tilde{Y}} f_{\eta}(\tilde{d}(z, y)) f_y(y) \frac{\partial \tilde{d}(z, y)}{\partial z} dy$$

for  $z \in Z_d$ . For  $z \in \tilde{Z} \setminus Z_d$ , the density of demand is zero. Both inverse supply and inverse demand depend on  $P(z)$ , although the dependence is implicit.

Among the set of smooth price functions that yield unique interior optima, an equilibrium price function must satisfy the condition that the density of supply equals the density of demand for all values of  $z \in \tilde{Z}$ . Making explicit the dependence of the supply and demand densities on  $P$ , this condition requires that  $Z_s = Z_d$  and that

$$(2.5) \quad \begin{aligned} & \int_{\tilde{X}} f_{\varepsilon}(\tilde{s}(z, x; P(z))) f_x(x) \frac{\partial \tilde{s}(z, x; P(z))}{\partial z} dx \\ &= \int_{\tilde{Y}} f_{\eta}(\tilde{d}(z, y; P(z))) f_y(y) \frac{\partial \tilde{d}(z, y; P(z))}{\partial z} dy \end{aligned}$$

for all  $z \in Z_s \cap Z_d$ , where  $P(z)$  is the price function. The equilibrium price function solves this equation.

### 2.1. Properties of Equilibrium

Ekeland (2010) and Chiappori, McCann, and Nesheim (2010) showed that an equilibrium exists in the hedonic model under very general conditions which include those given in the previous section. They also showed that the single-crossing property presented in the previous section is sufficient for equilibrium to be unique and pure in the sense that each worker matches to a single firm and each pair chooses a single job type  $z$ . This implies, in particular, that counterfactual policy analysis using this model provides unambiguous answers.

A key implication of the model is that cross-market variation in the price function is determined by cross-market variation in the functions  $(f_x, f_{\varepsilon})$

and  $(f_y, f_\eta)$ . To see this more clearly, substitute for  $\frac{\partial \tilde{s}(z, x)}{\partial z}$  and  $\frac{\partial \tilde{d}(z, y)}{\partial z}$ , and rewrite (2.5) as

$$(2.6) \quad P_{zz}(z) = \frac{\int_{\tilde{Y}} f_\eta f_y \frac{\Gamma_{zz}}{\Gamma_{z\eta}} dy - \int_{\tilde{X}} f_\varepsilon f_x \frac{U_{zz}}{U_{z\varepsilon}} dx}{\int_{\tilde{Y}} \frac{f_\eta f_y}{\Gamma_{z\eta}} dy - \int_{\tilde{X}} \frac{f_\varepsilon f_x}{U_{z\varepsilon}} dx},$$

where the arguments of the functions are suppressed. This second order differential equation depends on the objects  $(U, \Gamma)$ ,  $(f_x, f_\varepsilon)$ , and  $(f_y, f_\eta)$ . If the parameters  $(U, \Gamma)$  and  $(f_\varepsilon, f_\eta)$  are fixed while  $(f_x, f_y)$  vary across markets, the equilibrium price will vary across markets. We exploit the variation induced by these changes in Section 4 to show identification of the structural parameters using multimarket data.

Finally, our assumptions imply that the joint distribution of observables  $(z, x, y)$  implied by the model will be continuous. Since the distributions of  $(x, \varepsilon)$  and  $(y, \eta)$  are continuous, the joint distribution of  $(z, x, y)$  will be continuous if almost all agents do not bunch at a single point in  $\tilde{Z}$ . Given a point  $z_b \in \tilde{Z}$ , it is possible that a positive measure of agents will bunch at  $z_b$  only if the set

$$B = \{(x, \varepsilon), (y, \eta) \mid \Gamma_z(z_b, y, \eta) = U_z(z_b, x, \varepsilon)\}$$

has positive measure. When  $(x, \varepsilon)$  and  $(y, \eta)$  are continuously distributed, this set cannot have positive measure if  $\Gamma_{z\eta} > 0$  and  $U_{z\varepsilon} < 0$ . Our assumed single-crossing condition rules out bunching and implies that  $(z, x, y)$  are continuously distributed on their respective domains.<sup>11</sup>

### 3. IDENTIFICATION IN A SINGLE MARKET

This section discusses identification of the supply function  $s(x, \varepsilon)$ , the marginal utility function  $U_z(z, x, \varepsilon)$ , and the distribution of  $\varepsilon$  in a single market. We do not discuss the demand side of the market because the analysis for that case is analytically similar. Our analysis assumes that the equilibrium price function  $P(z)$  and the distribution of  $(z, x)$  are known, where  $z$  denotes the observed hedonic location choice of an individual and  $x$  denotes the vector of observed worker characteristics. Estimation of  $P(z)$  does not affect the identification analysis. In Section 5, we discuss how it impacts estimation. (See Technical Appendix D.)

<sup>11</sup>Examples with bunching can be generated either by allowing for mass points in the distributions of worker or firm types or by relaxing the single-crossing condition. Further analysis of bunching is available in Technical Appendix C.



We conduct the analysis for the case where  $z \in \mathbf{R}$ . The analysis applies equally to the case where  $z = g(a) \in \mathbf{R}$  and  $a \in \mathbf{R}^n$  subject to a normalization on  $g$ . For example, if  $z = \beta \cdot a$  and the first coordinate of  $\beta$ ,  $\beta_1 = 1$ , then the vector of parameters for  $\beta$  can be estimated from the price function since the equilibrium price has the form  $p = P(\beta \cdot a)$ . Under these conditions  $(dP/da_j)/(dP/da_1) = \beta_j$ . Once  $\beta$  has been identified, the analysis in this paper can be applied to  $z = \beta \cdot a$ .<sup>12</sup>

In nonadditive hedonic models, the supply function  $s(x, \varepsilon)$  is a nonseparable function of a vector of observables  $x$  and a scalar unobservable  $\varepsilon$ . While we assume that  $\varepsilon$  is independent of  $x$ , nonetheless, the nonadditive structure allows for arbitrary interactions between  $x$  and  $\varepsilon$  in producing outcomes. Furthermore, the theoretical structure of Section 2 implies that  $s$  is a strictly increasing function of  $\varepsilon$ . The function  $s(x, \varepsilon)$  can then be identified using results from Matzkin (1999, 2003), where it is shown that, using the independence between  $\varepsilon$  and  $x$ , and the strict monotonicity of  $s$  in  $\varepsilon$ , one has that for all values  $x$  of  $X$  and all values  $\varepsilon'$  of  $\varepsilon$ ,

$$\begin{aligned} F_\varepsilon(\varepsilon') &= \Pr(\varepsilon \leq \varepsilon') \\ &= \Pr(\varepsilon \leq \varepsilon' | X = x) \\ &= \Pr(s(X, \varepsilon) \leq s(x, \varepsilon') | X = x) \\ &= F_{Z|X=x}(s(x, \varepsilon')). \end{aligned}$$

The conditional distribution function,  $F_{Z|X=x}$ , of  $Z$  given  $X = x$  is strictly increasing when the density of  $\varepsilon$  is everywhere positive. In this case,

$$s(x, \varepsilon') = F_{Z|X=x}^{-1}(F_\varepsilon(\varepsilon')).$$

Note that  $s(x, \varepsilon)$  is a reduced form. It depends on the structural utility function and on the equilibrium price  $P(z)$ . Nevertheless, it can be used to predict partial equilibrium impacts of changes in  $x$  or  $\varepsilon$  on equilibrium outcomes. Moreover, identification of  $s(x, \varepsilon)$  is a first step required for identification of  $U_z$ .

Identification of  $s(x, \varepsilon)$  requires either a normalization of the function (e.g., fixing its values at a particular value of  $X$ ) or of the distribution of the unobserved  $\varepsilon$  (e.g., assuming that the distribution  $F_\varepsilon(\varepsilon')$  is known). If, for example, we normalize  $s$  so that at a value  $\bar{x}$  of  $X$  and all  $\varepsilon'$ ,

$$s(\bar{x}, \varepsilon') = \varepsilon',$$

then one gets from the above equations that for all  $\varepsilon'$ ,

$$F_\varepsilon(\varepsilon') = F_{Z|X=\bar{x}}(s(\bar{x}, \varepsilon')) = F_{Z|X=\bar{x}}(\varepsilon')$$

<sup>12</sup>Epplé and Sieg (1999) and Sieg, Smith, Banzhaf, and Walsh (2004) used this index structure in their analysis of parametric hedonic models.

and for all  $x$ ,

$$s(x, \varepsilon') = F_{Z|X=x}^{-1}(F_{Z|X=\bar{x}}(\varepsilon')).$$

See Matzkin (1999, 2003) for details.

Certain features of the function  $s$ , such as  $\partial s(x, \varepsilon')/\partial x$  (the partial derivative of  $s$  with respect to  $x$ ) or  $z_1 - z_0 = s(x_1, \varepsilon) - s(x_0, \varepsilon)$  (the change in  $z$  due to a change in  $x$  from  $x_0$  to  $x_1$ , leaving the value of  $\varepsilon$  fixed), are invariant to the choice of a normalization. These can be expressed as

$$\frac{\partial s(x, \varepsilon')}{\partial x} = - \left( \frac{\partial F_{Z|X=x}(z)}{\partial z} \right)^{-1} \left( \frac{\partial F_{Z|X=x}(z)}{\partial x} \right),$$

where  $z = s(x, \varepsilon')$ , and as

$$s(x_1, \varepsilon) - s(x_0, \varepsilon) = F_{Z|X=x_1}^{-1}(F_{Z|X=x_0}(z_0)) - z_0.$$

See Matzkin (2007) for details. We next consider assumptions under which  $U_z$  is identified. We start with a nonidentification result that motivates our identification analysis.

### 3.1. A Nonidentification Result

Given  $s(x, \varepsilon)$  and  $P(z)$ , we seek to identify the marginal utility function  $U_z$ . In equilibrium, this function satisfies the first order condition

$$(3.1) \quad U_z(s(x, \varepsilon), x, \varepsilon) = P_z(s(x, \varepsilon)).$$

Note that the marginal utility function is identified for those values of  $(z, x, \varepsilon)$  that lie on the surface  $\{(z, x, \varepsilon) | z = s(x, \varepsilon)\}$ . On this surface, the value of the marginal utility  $U_z$  is known, since it must equal the value of the marginal price function.

However, it is clear from this expression that without further restrictions, it is not possible to identify the function  $U_z$  for all values of  $(z, x, \varepsilon)$  using data from a single market. For any arbitrary values of  $x$  and  $\varepsilon$ , the value  $s(x, \varepsilon)$ , the first argument of the function, is uniquely determined. Thus, even if we could observe  $\varepsilon$ , we could not independently vary  $(z, x, \varepsilon)$  and trace out the function on its  $n_x + 2$  dimensional domain.<sup>13</sup>

We offer three responses to this fundamental nonidentification problem: (i) Focus attention on features of  $U_z$  that are identified; (ii) impose functional restrictions on  $U_z$  that enable analysts to overcome the exact functional dependence between  $z$ ,  $x$ , and  $\varepsilon$  that is implied by economic theory, or (iii) obtain

<sup>13</sup>This is the nonparametric generalization of the nonidentification result reported in Brown and Rosen (1982).

data from equilibria in different markets and make use of independent variation in hedonic equilibrium prices across markets. We pursue the first two approaches in the remainder of this section and develop the third approach in Section 4.

### 3.2. What Is Identified Without Further Structure

Even though  $U_z$  is not identified using data from a single market, some features of the function  $U_z$  are identified. A revealed preference argument shows that the hedonic price function provides a bound on  $U_z$ . If a worker  $(x_0, \varepsilon_0)$  chooses  $z_0$  and not  $z_1$ , where  $z_0 > z_1$ , then it must be the case that

$$\begin{aligned} P(z_0) - P(z_1) &\geq U(z_0, x_0, \varepsilon_0) - U(z_1, x_0, \varepsilon_0) \\ &= \int_{z_1}^{z_0} U_z(s, x_0, \varepsilon_0) ds. \end{aligned}$$

Differences in hedonic prices for  $z_0$  and  $z_1$  in general overestimate the change in welfare between  $z_1$  and  $z_0$ .<sup>14</sup>

Other features of  $U_z$  are point identified. For example, if  $x$  contains two variables,  $x_1$  and  $x_2$ , the ratio of the partial derivatives of  $U_z$  with respect to  $x_1$  and  $x_2$  is identified. To see this, note that one can totally differentiate equation (3.1) with respect to  $x_1$  and  $x_2$  to obtain

$$\begin{aligned} U_{zz}(z, x_1, x_2, \varepsilon) \frac{\partial s(x_1, x_2, \varepsilon)}{\partial x_1} + U_{zx_1}(z, x_1, x_2, \varepsilon) \\ = P_{zz}(s(x_1, x_2, \varepsilon)) \frac{\partial s(x_1, x_2, \varepsilon)}{\partial x_1} \end{aligned}$$

and

$$\begin{aligned} U_{zz}(z, x_1, x_2, \varepsilon) \frac{\partial s(x_1, x_2, \varepsilon)}{\partial x_2} + U_{zx_2}(z, x_1, x_2, \varepsilon) \\ = P_{zz}(s(x_1, x_2, \varepsilon)) \frac{\partial s(x_1, x_2, \varepsilon)}{\partial x_2}. \end{aligned}$$

Hence, from the assumed properties of  $U$  and the additional assumption that  $U_{zz} - P_{zz}$  is nonzero in equilibrium,

$$(3.2) \quad \frac{U_{zx_1}(z, x_1, x_2, \varepsilon)}{U_{zx_2}(z, x_1, x_2, \varepsilon)} = \frac{\frac{\partial s(x_1, x_2, \varepsilon)}{\partial x_1}}{\frac{\partial s(x_1, x_2, \varepsilon)}{\partial x_2}} \bigg|_{s(x_1, x_2, \varepsilon)=z}.$$

<sup>14</sup>See Scotchmer (1985), Kanemoto (1988), and Griffith and Nesheim (2008).

Since the derivatives of  $s(x, \varepsilon)$  are identified, the ratios of partial derivatives in (3.2) are identified without any further restrictions. The ratio on the left hand side of (3.2) measures the effect on  $U_z$  of changing  $x_1$  relative to changing  $x_2$ . Since  $U_z$  is the marginal willingness to accept (WTA) for the worker who chooses  $z$ , this ratio measures the relative impact on WTA of  $x_1$  and  $x_2$ . For example, when  $z$  is injury risk and  $x_1$  and  $x_2$  are education and experience, respectively, the ratio measures the relative impacts on WTA of education and experience. This identification result requires no further restrictions on the set of admissible  $U_z$  functions other than what has been assured thus far. It does not require any normalization of the  $s$  function.

Similarly, if  $x$  is scalar and one assumes that the distribution of  $\varepsilon$  is known, the same type of arguments can be used to show that the ratio of partial derivatives

$$(3.3) \quad \frac{U_{zx}(z, x, \varepsilon)}{U_{z\varepsilon}(z, x, \varepsilon)} = \frac{\frac{\partial s(x, \varepsilon)}{\partial x}}{\frac{\partial s(x, \varepsilon)}{\partial \varepsilon}} \bigg|_{s(x, \varepsilon)=z}$$

is identified without any further restrictions on  $U_z$ . In this case, (3.3) can be used to evaluate the relative impacts on  $U_z$  of observed  $x$  and unobserved  $\varepsilon$  for different values of  $x$  and at different quantiles of the distribution of  $\varepsilon$ . In the job injury example, this ratio could be used to study the impact of education on WTA for injury risk at different quantiles of the distribution of unobservables.

### 3.3. Identification Assuming Further Information About the Structure of Preferences and Technology

A second way to deal with the fundamental nonidentification problem is to assume additional restrictions on the set of admissible  $U_z$  functions. For example, Ekeland, Heckman, and Nesheim (2004) overcame the nonidentification problem by imposing an additive structure on  $U_z$ . In their leading example, they assumed that

$$U_z(z, x, \varepsilon) = u_0(z) - u_1(x) - \varepsilon.$$

This assumption implies that heterogeneity shifts the marginal utility of  $z$ , but has no effect on higher order derivatives. All workers have the same curvature in the utility function.

Under this restriction (and similar restrictions discussed in Ekeland, Heckman, and Nesheim (2004)), equation (3.1) reduces to

$$u_0(s(x, \varepsilon)) - u_1(x) - \varepsilon = P_z(s(x, \varepsilon)).$$

Ekeland, Heckman, and Nesheim (2004) showed that this implies that  $u_0$ ,  $u_1$ , and the distribution of  $\varepsilon$  are identified up to scale and location parameters.

However, the additive structure may be too strong for certain economic environments. From an economic perspective, it significantly limits the role of heterogeneity in hedonic models. Moreover, it imposes testable restrictions on observed data. Under the additive model restriction, the ratios analogous to (3.2) are

$$\frac{\frac{\partial u_1}{\partial x_1}(x_1, x_2)}{\frac{\partial u_1}{\partial x_2}(x_1, x_2)} = \frac{\frac{\partial s(x_1, x_2, \varepsilon)}{\partial x_1}}{\frac{\partial s(x_1, x_2, \varepsilon)}{\partial x_2}} \bigg|_{s(x_1, x_2, \varepsilon)=z}.$$

The rate of trade-off between  $x_1$  and  $x_2$  must be the same for all quantiles of the  $z$  distribution. This is a testable restriction. In the case of scalar  $x$ , we have

$$\frac{\partial u_1(x)}{\partial x} = \frac{\frac{\partial s(x, \varepsilon)}{\partial x}}{\frac{\partial s(x, \varepsilon)}{\partial \varepsilon}} \bigg|_{s(x, \varepsilon)=z}.$$

The impact of  $x$  on  $z$  must be the same for all  $\varepsilon$ .

The additive model structure is overidentified. We investigate weaker conditions for identification. Proceeding along these lines, we develop three theorems. Each introduces a different assumption on marginal utility, but all have the effect of reducing the number of arguments of  $U_z$  by at least one to overcome the nonidentification result of Section 3.1. All impose some separability structure on the utility or production functions and some sort of scale normalization. The types of restrictions we study are similar in spirit to either those required to identify household equivalence scales in classical demand theory or those used in random coefficient models. See Lewbel (1989, 1997) for related results on identification of household equivalence scales and a survey of that literature. For all the theorems below, we assume that the assumptions in Section 2 are satisfied. In particular, (i)  $U_z$  and  $P_z$  are continuously differentiable, (ii) for all  $\varepsilon \in \tilde{E}$  and  $x \in \tilde{X}$ , the supply function  $s(x, \varepsilon)$  satisfies (2.1) and (2.2), (iii)  $\varepsilon$  is distributed independently of  $x$  with a distribution,  $F_\varepsilon$ , that is continuously differentiable and strictly increasing on the compact set  $\tilde{E}$ , and (iv) the density  $f_x$  of  $X$  is strictly positive on the compact set  $\tilde{X}$ .

### 3.3.1. Observable Scales

For our first theorem, we assume that  $U_z$  is a weakly separable function in the pair  $(z, x_1)$ . We let  $(x_1, x_2) \in \tilde{X}_1 \times \tilde{X}_2 = \tilde{X}$ , where  $\tilde{X}_2$  may be the empty set. For some functions  $m: \mathbf{R} \times \tilde{X}_2 \times \tilde{E} \rightarrow \mathbf{R}$  and  $q: \tilde{Z} \times \tilde{X}_1 \rightarrow \mathbf{R}$ ,  $U_z(z, x_1, x_2, \varepsilon) = m(q(z, x_1), x_2, \varepsilon)$ . We assume that the function,  $q$ , which aggregates the effect of  $z$  and  $x_1$  on  $U_z$ , is known. In the job risk example,

$q(z, x_1)$  can be interpreted as the observable scale or quality of the job–worker match, while  $\varepsilon$  is unobserved heterogeneity in marginal utility. Note that when  $q(z, x_1) = z/x_1$ , our specification is in the same spirit as [Barten's \(1964\)](#) scale model for utility. Since  $m$  is nonadditive in  $\varepsilon$ , a normalization on either  $m$  or the distribution of  $\varepsilon$  is also needed. In the next theorem, we assume that the distribution of  $\varepsilon$  is specified a priori.<sup>15</sup>

**THEOREM 3.1:** *Suppose that for some unknown continuously differentiable function  $m: \mathbf{R} \times \tilde{X}_2 \times \tilde{E} \rightarrow \mathbf{R}$ , which is strictly decreasing in its last argument, and some known continuously differentiable function  $q: \tilde{Z} \times \tilde{X}_1 \rightarrow \mathbf{R}$ , the marginal utility function can be written*

$$(3.4) \quad U_z(z, x_1, x_2, \varepsilon) = m(q(z, x_1), x_2, \varepsilon).$$

*Further assume that the distribution,  $F_{\varepsilon}$ , of  $\varepsilon$  is known. For all  $(x_2, \varepsilon) \in \tilde{X}_2 \times \tilde{E}$ , let  $\tilde{Q}(x_2, \varepsilon) = \{t \in \mathbf{R} \mid \text{for some } x_1 \in \tilde{X}_1, q(s(x_1, x_2, \varepsilon), x_1) = t\}$ . Then, for all  $(z, x_1, x_2, \varepsilon)$  such that  $(x_2, \varepsilon) \in \tilde{X}_2 \times \tilde{E}$  and  $q(z, x_1) \in \tilde{Q}(x_2, \varepsilon)$ ,  $U_z(z, x_1, x_2, \varepsilon)$  is identified.*

See [Appendix A](#) for the proof.

The intuition for this result is that for a given value of  $(x_2, \varepsilon)$ , identification of  $s(x_1, x_2, \varepsilon)$  allows one to find all pairs of  $(z, x_1)$  that are consistent with that value of  $(x_2, \varepsilon)$ . The separability restriction on  $U_z$  allows one to select from among these pairs the one that produces a given value of  $q(z, x_1)$ . Combining these two results allows one to identify  $U_z$  at any arbitrary point.

[Theorem 3.1](#) uses a separability restriction and a normalization of the distribution of  $\varepsilon$ . An alternative approach is to invoke separability, but to normalize the function  $U_z$ , instead of the distribution of  $\varepsilon$ , by assuming that its value is known at all points on a one dimensional curve. Along these lines, we obtain the following theorem.

**THEOREM 3.2:** *Suppose that for some unknown continuously differentiable function  $m: \mathbf{R} \times \tilde{X}_2 \times \tilde{E} \rightarrow \mathbf{R}$ , which is strictly decreasing in its last argument, and some known continuously differentiable function  $q: \tilde{Z} \times \tilde{X}_1 \rightarrow \mathbf{R}$ ,*

$$U_z(z, x, \varepsilon) = m(q(z, x_1), x_2, \varepsilon).$$

*Let  $\bar{x} \in \tilde{X}$  be specified and normalize  $\tilde{Z} = \tilde{E}$ . Define the line  $\{(z, x, \varepsilon) \mid z = \varepsilon, x = \bar{x}\}$ . For all points on this line, use the function  $P_z$  to fix the value of the unknown function  $U_z$  by requiring that for all  $z \in \tilde{Z}$ ,*

$$(3.5) \quad U_z(z, \bar{x}, z) = P_z(z).$$

<sup>15</sup>For example, since  $\varepsilon$  is not observed, a natural choice is to specify that  $\varepsilon$  is uniformly distributed. Thus  $\varepsilon$  can be interpreted as the quantile in the unobserved heterogeneity distribution.

For all  $(x_2, \varepsilon) \in \tilde{X}_2 \times \tilde{E}$ , let  $\tilde{Q}(x_2, \varepsilon) = \{t \in \mathbf{R} \mid \text{for some } x_1 \in \tilde{X}_1, q(s(x_1, x_2, \varepsilon), x_1) = t\}$ . Then, for all  $(z, x_1, x_2, \varepsilon)$  such that  $(x_2, \varepsilon) \in \tilde{X}_2 \times \tilde{E}$  and  $q(z, x_1) \in \tilde{Q}(x_2, \varepsilon)$ ,  $U_z(z, x_1, x_2, \varepsilon)$  is identified.

See Appendix A for the proof.

### 3.3.2. Random Coefficients

Results analogous to Theorems 3.1 and 3.2 can be obtained if, instead of requiring that  $U_z$  be separable in  $(z, x_1)$ , one assumes that  $U_z$  is separable in  $(z, \varepsilon)$ . In other words, one can assume that for some *unknown* function  $m: \mathbf{R} \times \tilde{X} \rightarrow \mathbf{R}$ , which is strictly increasing in its first argument, and some *known* function  $q: \tilde{Z} \times \tilde{E} \rightarrow \mathbf{R}$ , which is strictly decreasing in its second argument,

$$(3.6) \quad U_z(z, x, \varepsilon) = m(q(z, \varepsilon), x),$$

where both  $m$  and  $q$  are assumed to be continuously differentiable to guarantee that  $U_z$  is continuously differentiable as assumed in Section 2. This specification is a generalization of a linear random coefficients model. When  $q(z, \varepsilon) = z/\varepsilon$  and  $\tilde{E} \subset \mathbf{R}_{++}$ , this specification is consistent with a linear random coefficients model for the marginal utility.<sup>16</sup>

A result analogous to that of Theorem 3.1 can be obtained when  $F_\varepsilon$  is known. For any  $x \in \tilde{X}$ , let  $\tilde{Q}(x) = \{t \in \mathbf{R} \mid \text{for some } \varepsilon \in \tilde{E}, q(s(x, \varepsilon), \varepsilon) = t\}$ . Following an argument similar to that used in the proof of Theorem 3.1, it is easy to show that for all  $(z, x, \varepsilon)$  such that  $(x, \varepsilon) \in \tilde{X} \times \tilde{E}$  and  $q(z, \varepsilon) \in \tilde{Q}(x)$ ,  $U_z(z, x, \varepsilon)$  is identified. The result that corresponds to Theorem 3.2 is obtained by fixing the values of  $m$  when  $x = \bar{x}$  and  $z = \varepsilon$ , as

$$m(q(z, z), \bar{x}) = P_z(z).$$

### 3.3.3. Multiple Observed Scales

In the previous cases, the vector  $x$  could possess only one coordinate. When  $x$  is a vector, alternative restrictions can be used to achieve identification. As an example, when  $x$  is a two-dimensional vector  $(x_1, x_2)$ , we can impose the restriction that  $U_z$  is weakly separable into two known functions,  $q_1(z, x_1)$  and  $q_2(x_2, \varepsilon)$ . The next theorem establishes that this restriction, along with a normalization on  $U_z$ , which is weaker than what is required in the case where  $x$  is a scalar, allows one to identify  $U_z$ .

<sup>16</sup>The scalar version of the parametric demand model analyzed in Bajari and Benkard (2005) is a special case of this model. They assumed that  $m$  is parametric.

**THEOREM 3.3:** Let  $x = (x_1, x_2, x_3) \in \tilde{X}_1 \times \tilde{X}_2 \times \tilde{X}_3 = \tilde{X}$ . Suppose that for some unknown continuously differentiable function  $m: \mathbf{R}^2 \times \tilde{X}_3 \rightarrow \mathbf{R}$ , which is strictly increasing in its second argument, and some known continuously differentiable functions  $q_1: \tilde{Z} \times \tilde{X}_1 \rightarrow \mathbf{R}$  and  $q_2: \tilde{X}_2 \times \tilde{E} \rightarrow \mathbf{R}$ .

$$(3.7) \quad U_z(z, x_1, x_2, x_3, \varepsilon) = m(q_1(z, x_1), q_2(x_2, \varepsilon), x_3)$$

where  $q_2$  is strictly decreasing in its second argument. Let  $[q_2^l, q_2^u]$  denote the support of  $q_2(x_2, \varepsilon)$ . Assume that the function  $U_z$  is known at one point  $(\bar{z}, \bar{x}_1, \bar{x}_2, \bar{x}_3, \bar{\varepsilon}) \in \tilde{Z} \times \tilde{X} \times \tilde{E}$ , so that

$$(3.8) \quad U_z(\bar{z}, \bar{x}_1, \bar{x}_2, \bar{x}_3, \bar{\varepsilon}) = P_z(\bar{z}).$$

Let  $\bar{\alpha} \in [q_2^l, q_2^u]$  be such that  $q_2(\bar{x}_2, \bar{\varepsilon}) = \bar{\alpha}$ , and assume that for all  $\varepsilon \in \tilde{E}$ , there exists  $x_2 \in \tilde{X}_2$  such that  $q_2(x_2, \varepsilon) = \bar{\alpha}$ . For all  $t_2 \in [q_2^l, q_2^u]$  and  $x_3 \in \tilde{X}_3$ , let  $\tilde{Q}(t_2, x_3) = \{t_1 \in \mathbf{R} \mid \text{for some } (x_1, x_2, \varepsilon) \in \tilde{X}_1 \times \tilde{X}_2 \times \tilde{E}, q_2(x_2, \varepsilon) = t_2 \text{ and } q_1(s(x_1, x_2, x_3, \varepsilon), x_1) = t_1\}$ . Then, for all  $(z, x_1, x_2, x_3, \varepsilon)$  such that  $(x_2, \varepsilon) \in \tilde{X}_2 \times \tilde{E}$  and  $q_1(z, x_1) \in \tilde{Q}(t_2, x_3)$ ,  $U_z(z, x_1, x_2, x_3, \varepsilon)$  is identified.

See Appendix A for the proof.

Identification of  $U_z$  is obtained from a sequence of arguments. First, equation (3.7) implies that the supply function  $s(x_1, x_2, x_3, \varepsilon)$  is a function of  $x_1, x_3$ , and  $q_2(x_2, \varepsilon)$ . Equation (3.8) implies that the supply function is known at one point. Further, the strict monotonicity of  $m$  and  $q_2$  in their second arguments implies that the supply function is strictly increasing in  $\varepsilon$ . These implications guarantee that the supply function  $s$  and the distribution of  $\varepsilon$  are identified. Next, to identify the value of  $m(t_1, t_2, x_3)$  at an arbitrary point  $(t_1, t_2, x_3)$  on the relevant domain, we first find values  $x_1^*, x_2^*$ , and  $\varepsilon^*$  such that  $z = s(x_1^*, x_2^*, x_3, \varepsilon^*)$ ,  $q_1(z, x_1^*) = t_1$ , and  $q_2(x_2^*, \varepsilon^*) = t_2$ . Finally, since such a  $z$  satisfies the first order condition (FOC), it follows that  $m(t_1, t_2, x_3) = P_z(z) = P_z(s(x_1^*, x_2^*, x_3, \varepsilon^*))$ . Thus, independent variation in  $x_1$  and  $x_2$ , the assumed dependence of  $U_z$  on only two arguments, and knowledge of the functions  $q_1$  and  $q_2$  allow one to trace out  $U_z$  as a function of its two arguments.

The statement and proof of Theorem 3.3 can easily be modified to show that the function  $U_z$  is also identified when it can be expressed as a function  $m(t_1, x_1, x_3)$ , where  $t_1 = q_1(z, t_2)$  and  $t_2 = q_2(x_2, \varepsilon)$ . To see this, suppose that for some unknown function  $m: \mathbf{R} \times \tilde{X}_1 \times \tilde{X}_3 \rightarrow \mathbf{R}$  and some known functions  $q_1: \tilde{Z} \times \mathbf{R} \rightarrow \mathbf{R}$  and  $q_2: \tilde{X}_2 \times \tilde{E} \rightarrow \mathbf{R}$ , such that  $m$  is strictly increasing in its first argument,  $q_1$  is strictly increasing in its second argument, and  $q_2$  is strictly decreasing in its second argument,

$$(3.9) \quad U_z(z, x_1, x_2, x_3, \varepsilon) = m(q_1(z, q_2(x_2, \varepsilon)), x_1, x_3).$$

Assume that the function  $m$  is known at one point  $(\bar{z}, \bar{x}_1, \bar{\alpha}, \bar{x}_3)$ , so that

$$(3.10) \quad m(q_1(\bar{z}, \bar{\alpha}), \bar{x}_1, \bar{x}_3) = P_z(\bar{z}).$$



Then, as in the proof of Theorem 3.3, it can be shown by using (3.9) that the supply function,  $s(x_1, x_2, x_3, \varepsilon)$ , is weakly separable in  $(x_2, \varepsilon)$ . By (3.10), the value of  $s$  is known at one point. By the monotonicity of  $m$  and  $q_1$ ,  $s$  is strictly increasing in  $\varepsilon$ . These properties guarantee identification of  $s$  and of the distribution of  $\varepsilon$  using the analysis of Matzkin (1999, 2003). To identify the value of  $m(t_1, t_2, x_3)$  at an arbitrary vector  $(t_1, t_2, x_3)$ , let  $x_1^* = t_2$ , and find  $x_2^*$ , and  $\varepsilon^*$  such that when  $z = s(x_1^*, x_2^*, x_3, \varepsilon^*)$ ,  $q_1(z, q_2(x_2^*, \varepsilon^*)) = t_1$ . Then, as in the previous argument,  $m(t_1, t_2, x_3) = P_z(z) = P_z(s(x_1^*, x_2^*, x_3, \varepsilon^*))$ .

### 3.3.4. Economic Implications of the Restrictions

Consider the economic implications of some of the restrictions that we have invoked. We focus on the restrictions used in Theorems 3.1 and 3.2 with  $x = x_1$ . Consider the alternative specifications

$$(3.11) \quad U_z(z, x, \varepsilon) = m_1(zx, \varepsilon)$$

and, for  $\tilde{E} \subset \mathbf{R}_{++}$ ,

$$(3.12) \quad U_z(z, x, \varepsilon) = m_2(z/\varepsilon, x).$$

These models make different assumptions about how heterogeneity impacts marginal utility. Model (3.11) imposes the requirement that  $U_{zz}/U_{zx} = x/z$ , while model (3.12) imposes the requirement that  $U_{zz}/U_{z\varepsilon} = -\varepsilon/z$ . The former condition might be appropriate when the researcher has prior information about the marginal rate of substitution between  $z$  and observable  $x$ , but has no information about how unobserved heterogeneity affects marginal utility. The latter might be appropriate if the researcher is willing to impose a random coefficients structure, but has no prior information about the effects of observables.

## 4. IDENTIFICATION IN MULTIPLE MARKETS

The possibilities for identification of  $U_z(z, x, \varepsilon)$  in a single market are limited because all workers face a unique price schedule. Across multiple markets, the marginal price function  $P_z(z)$  will typically vary depending on underlying market conditions. For example, assuming that the marginal utility function,  $U_z(z, x, \varepsilon)$ , does not vary across markets, the marginal price function (and the supply function  $s(x, \varepsilon)$ ) will, in general, vary across markets when the distributions of worker or firm attributes vary across markets. When data are available from multiple markets and cross-market variation in the distributions of observables causes cross-market variation in  $P_z(z)$  and  $s(x, \varepsilon)$ , such variation can be used to identify the function  $U_z(z, x, \varepsilon)$  without invoking the conditions of Section 3.

Our analysis provides a general approach to identification of  $U_z$  under weaker conditions than are required when using data from a single cross section. Use of multimarket data to identify hedonic models is proposed in [Brown and Rosen \(1982\)](#), [Brown \(1983\)](#), and [Epple \(1987\)](#). Our analysis is more general than theirs because we consider the nonseparable case, whereas their analyses assume linearity of supply and price equations. Also, our approach makes explicit that the equilibrium price and the supply function depend on the distributions of observable characteristics of firms and workers.

Suppose that the distributions of  $\varepsilon$  and  $\eta$  are the same in all markets. Further assume that the distributions of  $x$  and  $y$  with densities denoted by  $(f_x, f_y) \in \mathcal{F}(\tilde{X}) \times \mathcal{F}(\tilde{Y}) \subseteq L_2(\tilde{X}, \mu_x) \times L_2(\tilde{Y}, \mu_y)$  vary across markets. Here,  $\mu_x$  and  $\mu_y$  are Lebesgue measures on  $\tilde{X}$  and  $\tilde{Y}$ , respectively, and  $L_2$  represents the space of square integrable functions. Suppose that a multimarket sample exists from  $M$  markets. The marginal price and supply functions in market  $m$  will depend on  $(f_x^m, f_y^m)$ , the densities of observed  $x$  and  $y$  in that market. Dropping superscripts, write these functions as  $P_z(z, f_x, f_y)$  and  $s(x, \varepsilon, f_x, f_y)$ .

The functions  $(f_x, f_y)$  and the functional  $P_z(z, f_x, f_y)$  are identified in a multimarket sample satisfying these conditions. Using results in [Matzkin \(1999, 2003\)](#), the functional  $s(x, \varepsilon, f_x, f_y)$  is nonparametrically identified. Hence, multimarket data allow one to identify the distributions of the observables and the dependence of the marginal price and the supply function on these distributions. We use this information to identify the marginal utility function  $U_z(z, x, \varepsilon)$  in the following way.

From the workers' first order condition,

$$U_z(s(x, \varepsilon, f_x, f_y), x, \varepsilon) = P_z(s(x, \varepsilon, f_x, f_y), f_x, f_y).$$

We have made explicit the dependence of  $P_z$  and  $s$  on  $f_x$  and  $f_y$ . In a single cross section, the price function and the supply function are fixed, and we cannot independently vary the three arguments of  $U_z$ . With multimarket data, both  $P_z$  and  $s$  may vary for each  $(x, \varepsilon)$  provided that  $f_x, f_y$ , or both vary across markets.

Before stating and proving [Theorem 4.1](#), we need to introduce some new notation. For each  $(f_x^m, f_y^m) \in \mathcal{F}(\tilde{X}) \times \mathcal{F}(\tilde{Y})$ , let  $s^m(x, \varepsilon)$  and  $P^m(z)$  be, respectively, the supply function and price function implied by equilibrium condition [\(2.5\)](#). Further, let

$$(4.1) \quad S = \{(z, x, \varepsilon) | z = s^m(x, \varepsilon) \text{ for some } (f_x^m, f_y^m) \in \mathcal{F}(\tilde{X}) \times \mathcal{F}(\tilde{Y})\}.$$

Similarly, let  $d^m(y, \eta)$  be the equilibrium demand function and

$$(4.2) \quad D = \{(z, y, \eta) | z = d^m(y, \eta) \text{ for some } (f_x^m, f_y^m) \in \mathcal{F}(\tilde{X}) \times \mathcal{F}(\tilde{Y})\}.$$

The sets  $S$  and  $D$  are the sets of matches that could occur in some feasible equilibria.

**THEOREM 4.1:** *Let  $S$  and  $D$  be defined as in (4.1) and (4.2). Assume that  $(f_\varepsilon, f_\eta)$  are known and are the same across markets. Then  $U_z(z, x, \varepsilon)$  is identified for all  $(z, x, \varepsilon) \in S$  and  $\Gamma_z(z, y, \eta)$  is identified for all  $(z, y, \eta) \in D$ .*

**PROOF:** Let  $(z, x, \varepsilon) \in S$ . Then  $z = s^m(x, \varepsilon)$  for some market  $m$  with  $(f_x^m, f_y^m) \subseteq \mathcal{F}(\tilde{X}) \times \mathcal{F}(\tilde{Y})$  and  $U_z(z, x, \varepsilon) = P_z^m(s^m(x, \varepsilon))$ , where  $P_z^m(z)$  is the marginal price in market  $m$ . Thus,  $U_z$  is identified at all points  $(z, x, \varepsilon)$  such that  $z$  is an equilibrium choice for  $(x, \varepsilon)$  in some feasible equilibrium. By an identical argument,  $\Gamma_z(z, y, \eta)$  is identified for all  $(z, y, \eta) \in D$ . *Q.E.D.*

Theorem 4.1 exploits the variation in price and supply functions induced by cross-market variation in the distributions of observables. Given a fixed pair  $(f_x^0, f_y^0)$ , under the conditions stated in Section 2, equation (2.5) will have a unique solution  $P^0(z)$ . This solution will imply a pair of supply and demand functions  $s^0(x, \varepsilon)$  and  $d^0(y, \eta)$ . Perturbing the underlying distributions to  $(f_x^1, f_y^1)$  will result in a new equilibrium price function  $P^1(z)$  and new supply and demand functions  $s^1(x, \varepsilon)$  and  $d^1(y, \eta)$ . In general, if the dependence of utility and production functions on  $y$  and  $x$  is not degenerate, it will be true that  $P^1(z) \neq P^0(z)$ ,  $s^1(x, \varepsilon) \neq s^0(x, \varepsilon)$ , and  $d^1(y, \eta) \neq d^0(y, \eta)$ .

## 5. ESTIMATION

Since our identification results are constructive, they suggest estimators for the marginal utility (resp. production) functions and their distributions, based on estimators for the derivative of the price function and the supply (resp. demand) function. The derivative of the price function can be estimated as the derivative of the kernel regression of price on  $z$ . The demand and supply functions, and the distributions of the unobserved variables, can be estimated also using kernel methods as described in Matzkin (2003).

The innovation in this paper over previous research is the estimation of the marginal utility and marginal product functions. Following our proofs of identification, we can construct estimators for either the supply or demand side. Consider, for example, the specification of the marginal utility in Theorem 3.3, where for some unknown function  $m$  and known functions  $q_1$  and  $q_2$ ,

$$U_z(z, x_1, x_2, x_3, \varepsilon) = m(q_1(z, x_1), q_2(x_2, \varepsilon), x_3).$$

The supply function resulting from this specification is separable in  $(x_2, \varepsilon)$ . An estimator for such supply function, denoted by  $\hat{s}(x_1, q_2(x_2, \varepsilon), x_3)$ , can be derived following Matzkin (2003). Let  $\hat{P}_z(z)$  denote the kernel estimator for the derivative of the price function, and let  $t_1$  and  $t_2$  be arbitrary values of  $q_1(\cdot)$  and  $q_2(\cdot)$ . It follows from the proof of Theorem 3.3 that

$$(5.1) \quad \hat{m}(t_1, t_2, x_3) = \hat{P}_z(\hat{s}(\hat{x}_1^*, t_2, x_3)),$$

where  $\hat{x}_1^*$  is the value of  $x_1$  that solves the equation

$$(5.2) \quad q_1(\hat{s}(x_1, t_2, x_3), x_1) = t_1.$$

Equation (5.1) with the estimated  $\hat{P}_z$  and  $\hat{s}$  plugged in is our estimator of the marginal utility function in this model.

Let  $x_1^*$  denote the value of  $x_1$  that satisfies

$$q_1(s(x_1^*, t_2, x_3), x_1^*) = t_1.$$

We demonstrate in Technical Appendix D that, under some conditions, the rate of convergence of  $\hat{x}_1^*$  to  $x_1^*$  is the same as that of  $\hat{s}$  to  $s$ . We develop the results using the Implicit Function Theorem in Banach spaces of [Hildebrandt and Graves \(1927\)](#) to derive a first order Taylor type approximation for the functional that maps the estimated distribution of the observable variables into the value of  $\hat{x}_1^*$ . Note that because  $q_1$  is known and  $(t_1, t_2)$  is given, equation (5.2) implies that  $\hat{s}(\hat{x}_1^*, t_2, x_3)$  is a deterministic function of  $\hat{x}_1^*$ . Hence, the asymptotic behavior of  $\hat{s}(\hat{x}_1^*, t_2, x_3)$  can be derived from that of  $\hat{x}_1^*$  using well known techniques. Also note that when the dimension of the vector  $(x_1, x_2, x_3)$  is strictly larger than the dimension of  $z$  plus 2, we may assume for the purpose of analyzing the asymptotic behavior of  $\hat{P}_z(\hat{s}(\hat{x}_1^*, t_2, x_3))$  that  $P_z$  is known. Hence, under standard regularity conditions,  $\hat{m}(t_1, t_2, x_3)$  will converge to  $m(t_1, t_2, x_3)$  at the same rate that  $\hat{s}$  converges to  $s$ , with an asymptotic variance that can be calculated from that of  $\hat{x}_1^*$  using the standard Delta method. See Technical Appendix D for additional detail and Technical Appendix E for some simulations that suggest that the method is practical and works well in practice.

## 6. CONCLUSIONS

This paper develops methods for identifying the structural parameters of hedonic equilibrium models where both the marginal utility of workers and the marginal product of firms are nonadditive functions of attributes and a random vector of individual characteristics, which are different for the workers and firms. We develop sufficient conditions to identify the marginal utility and marginal product functions using both single market and multimarket data. In the single market case, we develop nonparametric estimators for these functions and show in Technical Appendix D that they are consistent and asymptotically normal. Limited Monte Carlo evidence presented in Technical Appendix E suggests that the methods work well in practice.

Our analysis is for an hedonic model with a single attribute. However, as discussed in Sections 2 and 3, these results apply equally to a model in which the hedonic attribute is a single index with some known structure in terms of a vector of characteristics. We also note that our methods allow one to deal with unobserved product attributes, as long as these can be identified from the price function, as in [Bajari and Benkard \(2005\)](#), since in such a case one

can treat them as observed. In contrast to their analysis, our analysis allows for a nonparametric utility function. More generally, the first order conditions in unrestricted multiple attribute hedonic models are systems of simultaneous equations. General conditions for nonparametric identification of systems of equations are given in [Matzkin \(2008\)](#). Further work is required to investigate which types of restrictions developed in [Matzkin \(2008\)](#) are most appropriate for hedonic applications.

## APPENDIX A: PROOFS

**PROOF OF THEOREM 3.1:** Since model (3.4) satisfies the conditions in Section 2,  $s(x, \varepsilon)$  is characterized by (2.1) and (2.2). Thus,  $s$  is nonadditive and strictly increasing in  $\varepsilon$ . Since  $\varepsilon$  is independent of  $X$  and  $F_{Z|X=x}$  is strictly increasing at  $z = s(x, \varepsilon)$ , it follows by [Matzkin \(1999, 2003\)](#) that

$$s(x, \varepsilon) = F_{Z|X=x}^{-1}(F_{\varepsilon}(\varepsilon)).$$

Since  $F_{\varepsilon}$  is known,  $s(x, \varepsilon)$  is identified. Let  $(t_1, x_2, \varepsilon)$  be such that  $(x_2, \varepsilon) \in \tilde{X}_2 \times \tilde{E}$  and  $t_1 \in \tilde{Q}(x_2, \varepsilon)$ . By the definition of  $\tilde{Q}(x_2, \varepsilon)$ , there exists  $x_1^* \in \tilde{X}_1$  such that

$$q(s(x_1^*, x_2, \varepsilon), x_1^*) = t_1.$$

Then

$$m(t_1, x_2, \varepsilon) = P_z(s(x_1^*, x_2, \varepsilon)).$$

It follows that for all  $(z, x, \varepsilon)$  such that  $q(z, x_1) = t_1 \in \tilde{Q}(x_2, \varepsilon)$ ,

$$U_z(z, x, \varepsilon) = m(q(z, x_1), x_2, \varepsilon) = P_z(s(x_1^*, x_2, \varepsilon)). \quad Q.E.D.$$

**PROOF OF THEOREM 3.2:** As in Theorem 3.1,  $s$  is a nonadditive, strictly increasing function of  $\varepsilon$ . Further, it follows from equation (3.5) that the supply function,  $s(x, \varepsilon)$ , satisfies

$$s(\bar{x}, \varepsilon) = \varepsilon.$$

Then, by [Matzkin \(1999\)](#), for all  $\varepsilon \in \tilde{E}$ ,

$$F_{\varepsilon}(\varepsilon) = F_{Z|X=\bar{x}}(\varepsilon)$$

and, for all  $x \in \tilde{X}$ ,

$$s(x, \varepsilon) = F_{Z|X=x}^{-1}(F_{Z|X=\bar{x}}(\varepsilon)).$$

To establish that the function  $m$  is identified, let  $(t_1, x_2, \varepsilon)$  be such that  $(x_2, \varepsilon) \in \tilde{X}_2 \times \tilde{E}$  and  $t_1 \in \tilde{Q}(x_2, \varepsilon)$ . By the definition of  $\tilde{Q}(x_2, \varepsilon)$ , there exists  $x_1^* \in \tilde{X}_1$  such that

$$q(s(x_1^*, x_2, \varepsilon), x_1^*) = t_1.$$

Then,

$$m(t_1, x_2, \varepsilon) = P_z(s(x_1^*, x_2, \varepsilon)).$$

It follows that for all  $(z, x, \varepsilon)$  such that  $q(z, x_1) = t_1 \in \tilde{Q}(x_2, \varepsilon)$ ,

$$U_z(z, x, \varepsilon) = m(q(z, x_1), x_2, \varepsilon) = P_z(s(x_1^*, x_2, \varepsilon)). \quad Q.E.D.$$

**PROOF OF THEOREM 3.3:** Since  $U_z$  is weakly separable in  $(x_2, \varepsilon)$ , the function  $z = s(x_1, x_2, x_3, \varepsilon)$  is also weakly separable in  $(x_2, \varepsilon)$ . Hence, for some unknown function  $v$ ,

$$s(x_1, x_2, x_3, \varepsilon) = v(x_1, q_2(x_2, \varepsilon), x_3).$$

Further, since all of the conditions in Section 2 are satisfied,  $s$  is strictly increasing in  $\varepsilon$ . Since  $q_2$  is strictly decreasing in  $\varepsilon$ , then  $v$  is strictly decreasing in its second argument.

Let  $\bar{x}_2$  and  $\bar{\varepsilon}$  be such that  $q_2(\bar{x}_2, \bar{\varepsilon}) = \bar{\alpha}$ . Then, by separability and condition (3.8) in the statement of the theorem,

$$U_z(\bar{z}, \bar{x}_1, \bar{x}_2, \bar{x}_3, \bar{\varepsilon}) = P_z(\bar{z}).$$

Hence,  $\bar{z}$  satisfies the FOC (2.1) when  $x_1 = \bar{x}_1$  and  $q_2(\bar{x}_2, \bar{\varepsilon}) = \bar{\alpha}$ . It follows that

$$s(\bar{x}_1, \bar{x}_2, \bar{x}_3, \bar{\varepsilon}) = v(\bar{x}_1, q_2(\bar{x}_2, \bar{\varepsilon}), \bar{x}_3) = \bar{z}.$$

It then follows from the analysis of Matzkin (2003) that the function  $v$  and the distribution of  $\varepsilon$  are identified from the conditional distribution of  $z$  given  $(x_1, x_2, x_3)$ .

To show that the function  $m$  is identified, let  $(t_1, t_2, x_3)$  be any vector such that  $t_2 \in [q_2^l, q_2^u]$ ,  $x_3 \in \tilde{X}_3$ , and  $t_1 \in \tilde{Q}(t_2, x_3)$ . Let  $x_1^*$  denote a solution to

$$q_1(v(x_1^*, t_2, x_3), x_1^*) = t_1.$$

Since  $q_1$  is a known function and  $v$  can be recovered from the conditional cumulative distribution function of  $z$  given  $(x_1, x_2, x_3)$ , the only unknown in the expression is  $x_1^*$ . Since  $t_2 \in [q_2^l, q_2^u]$ ,  $x_3 \in \tilde{X}_3$ , and  $t_1 \in \tilde{Q}(t_2, x_3)$ ,  $x_1^*$  exists. Since  $v(x_1^*, t_2, x_3)$  satisfies the FOC (2.1),

$$\begin{aligned} \text{(A.1)} \quad m(t_1, t_2, x_3) &= m(q_1(v(x_1^*, t_2, x_3), x_1^*), t_2, x_3) \\ &= P_z(v(x_1^*, t_2, x_3)) \\ &= P_z(s(x_1^*, x_2^*, x_3, \varepsilon^*)) \end{aligned}$$

for any  $x_2^*$  and  $\varepsilon^*$  such that  $q_2(x_2^*, \varepsilon^*) = t_2$ . In (A.1), the first equality follows because  $q_1(v(x_1^*, t_2, x_3), x_1^*) = t_1$ . The second equality follows because the value of the marginal utility function  $m$  equals the value of the marginal price function at the particular value of  $z$  that satisfies the first order conditions. The third equality follows by the restriction on the function  $s$ . Since the function  $P_z$  is known and the function  $s$  can be recovered without knowledge of  $m$ , (A.1) implies that  $m$  is identified, and hence  $U_z$  is identified. *Q.E.D.*

## REFERENCES

- BAJARI, P., AND C. L. BENKARD (2005): "Demand Estimation With Heterogeneous Consumers and Unobserved Product Characteristics: A Hedonic Approach," *Journal of Political Economy*, 113, 1239–1276. [1570,1582,1587]
- BARTEN, A. P. (1964): "Family Composition, Prices and Expenditure Patterns," in *Econometric Analysis for National Economic Planning*, ed. by P. E. Hart, G. Mills, and J. K. Whitaker. London: Butterworths, 277–292. [1581]
- BERRY, S., AND A. PAKES (2007): "The Pure Characteristics Demand Model\*," *International Economic Review*, 48, 1193–1225. [1570]
- BROWN, J. N. (1983): "Structural Estimation in Implicit Markets," in *The Measurement of Labor Cost*, ed. by J. E. Triplett. Chicago: University of Chicago Press for NBER. [1570,1585]
- BROWN, J. N., AND H. S. ROSEN (1982): "On the Estimation of Structural Hedonic Price Models," *Econometrica*, 50, 765–768. [1570,1571,1577,1585]
- CHIAPPORI, P.-A., R. MCCANN, AND L. NESHEIM (2010): "Hedonic Price Equilibria, Stable Matching, and Optimal Transport: Equivalence, Topology, and Uniqueness," *Economic Theory*, 42, 317–354. [1573,1574]
- EKELAND, I. (2010): "Existence, Uniqueness and Efficiency of Equilibrium in Hedonic Markets With Multidimensional Types," *Economic Theory*, 42, 275–315. [1573,1574]
- EKELAND, I., J. J. HECKMAN, AND L. NESHEIM (2004): "Identification and Estimation of Hedonic Models," *Journal of Political Economy*, 112, S60–S109. [1570,1573,1579]
- EPPLÉ, D. (1987): "Hedonic Prices and Implicit Markets: Estimating Demand and Supply Functions for Differentiated Products," *Journal of Political Economy*, 95, 59–80. [1570,1571,1585]
- EPPLÉ, D., AND H. SIEG (1999): "Estimating Equilibrium Models of Local Jurisdictions," *Journal of Political Economy*, 107, 645–681. [1571,1576]
- GALE, D., AND H. NIKAIIDO (1965): "The Jacobian Matrix and Global Univalence of Mappings," *Mathematische Annalen*, 159, 81–93. [1573]
- GRIFFITH, R., AND L. NESHEIM (2008): "Household Willingness to Pay for Organic Products," Working Paper CWP18/08, CEMMAP. [1578]
- HECKMAN, J. J., AND E. J. VYTLACIL (2007): "Econometric Evaluation of Social Programs, Part I: Causal Models, Structural Models and Econometric Policy Evaluation," in *Handbook of Econometrics*, Vol. 6B, ed. by J. Heckman and E. Leamer. Amsterdam: Elsevier, 4779–4874. [1570]
- HECKMAN, J. J., R. L. MATZKIN, AND L. NESHEIM (2010): "Supplement to 'Nonparametric Identification and Estimation of Nonadditive Hedonic Models': Technical Appendix," *Econometrica Supplemental Material*, 78, [http://www.econometricsociety.org/ecta/Supmat/6388\\_extensions.pdf](http://www.econometricsociety.org/ecta/Supmat/6388_extensions.pdf). [1571]
- HILDEBRANDT, T. H., AND L. M. GRAVES (1927): "Implicit Functions and Their Differentials in General Analysis," *Transactions of the American Mathematical Society*, 29, 127–153. [1587]
- HURWICZ, L. (1962): "On the Structural Form of Interdependent Systems," in *Logic, Methodology and Philosophy of Science*, ed. by E. Nagel, P. Suppes, and A. Tarski. Stanford, CA: Stanford University Press, 232–239. [1570]
- KAHN, S., AND K. LANG (1988): "Efficient Estimation of Structural Hedonic Systems," *International Economic Review*, 29, 157–166. [1571]

- KANEMOTO, Y. (1988): "Hedonic Prices and the Benefits of Public Projects," *Econometrica*, 56, 981–989. [1578]
- KNIESNER, T. J., AND J. D. LEETH (1995): *Simulating Workplace Safety Policy*. Boston, MA: Kluwer Academic. [1572]
- LEETH, J. D., AND J. RUSER (2003): "Compensating Wage Differentials for Fatal and Nonfatal Risk by Gender and Race," *Journal of Risk and Uncertainty*, 27, 257–277. [1572]
- LEWBEL, A. (1989): "Identification and Estimation of Equivalence Scales Under Weak Separability," *Review of Economic Studies*, 56, 311–316. [1580]
- (1997): "Consumer Demand Systems and Household Equivalence Scales," in *Handbook of Applied Econometrics*, Vol. 2. Oxford, U.K.: Blackwell, 167–201. [1580]
- MATZKIN, R. L. (1999): "Nonparametric Estimation of Nonadditive Random Functions," Unpublished Manuscript, Department of Economics, Northwestern University. [1576,1577,1584,1585,1588]
- (2003): "Nonparametric Estimation of Nonadditive Random Functions," *Econometrica*, 71, 1339–1375. [1576,1577,1584–1586,1588,1589]
- (2007): "Nonparametric Identification," in *Handbook of Econometrics*, Vol. 6B, ed. by J. J. Heckman and E. Leamer. Amsterdam: Elsevier. [1577]
- (2008): "Identification in Nonparametric Simultaneous Equations Models," *Econometrica*, 76, 945–978. [1588]
- ROSEN, S. (1974): "Hedonic Prices and Implicit Markets: Product Differentiation in Pure Competition," *Journal of Political Economy*, 82, 34–55. [1570,1571]
- SCOTCHMER, S. (1985): "Hedonic Prices and Cost–Benefit Analysis," *Journal of Economic Theory*, 37, 55–75. [1578]
- SIEG, H., V. K. SMITH, H. S. BANZHAF, AND R. WALSH (2004): "Estimating the General Equilibrium Benefits of Large Changes in Spatially Delineated Public Goods," *International Economic Review*, 45, 1047–1077. [1571,1576]

*Dept. of Economics, University of Chicago, 1126 E. 59th Street, Chicago, IL 60637, U.S.A., and University College Dublin, American Bar Foundation and Cowles Foundation, Yale University; [jjh@uchicago.edu](mailto:jjh@uchicago.edu),*

*Dept. of Economics, University of California, Los Angeles, 8283 Bunche Hall, Los Angeles, CA 90095, U.S.A.; [matzkin@econ.ucla.edu](mailto:matzkin@econ.ucla.edu),*

*and*

*Dept. of Economics, University College London, Gower Street, London, WC1E 6BT, U.K. and CEMMAP and Institute for Fiscal Studies; [l.nesheim@ucl.ac.uk](mailto:l.nesheim@ucl.ac.uk).*

*Manuscript received March, 2006; final revision received April, 2009.*